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# Regression Models for Spatial Data: an Example from Gross Domestic Regional Bruto in Central Java Province

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#### Abstract

The important role of a region's transportation infrastructure strongly affects the economic growth of the region and tends to affect the surrounding areas. The effect is called spillover effect. The aim of the research was to recognize the direct effect and spillover effect (indirect) of transportation infrastructure on the economic growth in Central Java. To identify the spillover effects, it is necessary to recognize the different characteristics of each region which have the implications on the various transportation infrastructures at each region in Central Java. Therefore, the spatial modeling was conducted. In this study, the spatial modeling employed was Spatial Durbin Error Model (SDEM). The SDEM is another form of Spatial Error Model (SEM). It does not allow for lag effects of endogenous variables, but it allows for spatial error and spatial lag on exogenous variables in which it simplifies the interpretations on direct effects and spillover effect. According to SDEM estimates, the transportation infrastructures at the districts/municipalities in Central Java had no significant effect on the outputs at each region where the infrastructures were located and their neighboring districts/cities.

**Keywords:** Spatial Durbin Error Model, Spatial Error Model, Spatial Econometrics, Transportation Infrastructure, Economic Growth

JEL Classification: R11, R12, F43, C21

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## 1. Introduction

Central Java province is one of the provinces in Java. The economic growth of Central Java in 2014 (5.42 %), which was indicated by the growth rate of Gross Regional Domestic Product (GRDP) at the constant price of 2010, was higher than the previous year (5.14 % in 2013). It is quite reasonable given the relatively stable economic conditions since the onset of the global crisis in 2014 (Statistik, B.P, 2015). Figure 1 shows the trend of the economic growth in Central Java compared with the national economic growth. The economic growth in Central Java has continued to increase since 2013. In 2012 and

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2013, Indonesia's economic growth was at 5.34 % and 5.14 % respectively, while there was an upward trend in 2014 of 5.42 percent and 5.44 percent in 2015 (Statistik, B.P, 2016). In contrast to Indonesia's economic growth that declined in the last 3 years, the economic growth of Central Java was in an upward trend and higher than the national economic growth.

A relational approach to understanding differing economic geographies indicates that network systems are a key component of the regional development mix (Huggins & Thompson, 2017). The increase in the connectivity of districts and municipalities in Central Java has the impact on the productivity capability of each region to grow its economy. The increased rate of economic growth described earlier in this paper indicates that there is an increase in production capacity and economic process in Central Java. The increase in the number of products, goods and services also means that there is an increase in the number of the production factors used.

The existence of the different characteristic or phenomena structures owned by districts and cities such as the differences in capital value, number of labor, human capital, and transportation infrastructure can cause the variation of economic growth in each region. Moreover, it does not mean that each region is independent or unaffected by one another. All these consequences do not only occur in the areas where transport infrastructures are located, but they may also spread to neighboring areas. Increased inter-regional connectivity and/ or geographic proximity as well as characteristic similarity are the indicators of spillover effect.

The development of transportation infrastructure should be able to encourage economic growth in Central Java. In line with the increase in population demography, it is expected to encourage an increase in the use of transportation infrastructures. An important issue in this context is the influence of transportation infrastructure in meeting the growing demand of the people, so that it is required to meet the transportation needs in Central Java Province. Good regional connectivity is the goal in meeting the demand in the community so as to reduce capital, time and energy, because transportation is the main sector and the driver for economic growth in Central Java. The differences of regional characteristics have the implications on varied regional transportation infrastructure. Therefore, to see the effect of regional data-based





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transportation infrastructure with the approach of spatial durbin error model is a smart solution in solving the problems of the above phenomenon.

The first law on geography was proposed by Tobler (Ho & Hensher, 2016), which is the condition at one point or an area in relation with the condition of one of the points or adjacent areas. This law is the foundation for regional scientific studies, and the observations at one location frequently depend on the observations in other locations. To overcome this, the regression model is developed using spatial econometrics approach. Spatial effects often occur between one and another region.in spatial data, the observations in other neighboring locations.

The model of spatial econometrics has become increasingly popular in the applied econometrics of the last two decades, and the challenge and evolution in thinking of the best way to incorporate spatial heterogeneity in econometric models have emerged with this popularity (Lacombe & McIntyre, 2016). In spatial econometrics modeling, there are spatial autoregressive model (SAR) and spatial error model (SEM) as the special case of autoregressive spatial indicated by the addition of the effect of endogenous and exogenous lag variable known as spatial durbin model (SDM). In SDM, it has spatial lag in exogenous and endogenous variables, where  $\beta_2$  is the lag spatial parameter in  $W_1X$ . One alternative to SEM model is spatial durbin error model (SDEM), which does not allow for the effect of endogenous lag variable, but it allows for spatial error and spatial lag in exogenous variables. SDEM simplifies the interpretation of the direct impacts represented by the parameter model  $\beta$  and the indirect impact on y. SDEM replaces the global multiplier found in SDM, with local multiplier that makes it easier to interpret from model estimates (Lesage, 2008) (Karim, A & Wasono, R, 2017).

A review on the impact of transport infrastructure on economic growth had been undertaken by Banerjee, Duflo, & Qian (2012) which examined the relationship of the transport infrastructure access and economic growth in China by considering variables of inter-city distance using regression analysis through the Cobb-Douglas methodology. Furthermore, Arbués, Baños, & Mayor (2015) used spatial durbin model (SDM) found road transport infrastructure positively affects the output of the region in which the infrastructure is located and its neighboring provinces, while the remaining modes of transportation projects cause no significant impacts on average. Li, Wen, & Jiang (2017) also focused on the impacts of the transport infrastructure on regional economic growth in Chinese New Silk Road Economic Belt using SDM, obtaining the transport infrastructure plays an obvious lead role in regional economy growth alongside the NSREB, and the economic growth invigorates common development in surrounding regions. In addition, differences were observed among the different transport infrastructure with regard to their influences on regional economic development, as the highway transport affects regional economic growth to a larger degree than railway transport.

The primary goal of this paper is to estimate the contribution of transportation infrastructure to the regional economy by incorporating spatial effects and utilizing spatial econometrics techniques. The resulting model is applied to Province Central Java, where transportation infrastructure have been promoted through the implementation of the Infrastructure and Transport Strategic Plan that raised the quality of Province Central Java transportation network to National standards in a short period of time.

#### 2. Literature Review

#### 2.1 Regression Analysis

Regression analysis is one of the statistical methods that study the pattern of relationship mathematically between one dependent variable and one or more independent variables. The goal is to find out the value of dependent variable on the basis of the effect of independent variables.

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Mathematically, (Faraway, 2016) wrote as follows:

$$\mathbf{Y}_{t} = \boldsymbol{\beta}_{0} + \boldsymbol{\beta}_{1} \mathbf{X}_{1t} + \boldsymbol{\varepsilon}_{t}$$
(1)

where:

t	=	1, 2,, R		
y <sub>t</sub>	=	dependent variable		
$\mathbf{X}_{1t}, \mathbf{X}_{2t}, \dots, \mathbf{X}_{kt}$	=	independent variable with		
		the nature of permanent		
k	=	number of independent		
		variables		
$\beta_0, \beta_1,$	=	regression parameters		
$\mathcal{E}_t$	=	error (difference between		
		dependent variable and		
		regression model assessment)		

If the linear regression model in (1) is expressed in the form of the equation  $Y = X\beta + \varepsilon$ where Y and  $\varepsilon$  are the vectors with the dimension  $R \times 1$ , and X is a matrix with the size  $R \times (k + 1)$ , and  $\beta$  is a vector with the dimension  $(k + 1) \times 1$ , then using OLS estimates (Casela, Feinberg and Olkin, 2009):

$$\hat{\boldsymbol{\beta}} = \left(\mathbf{X}^T \mathbf{X}\right)^{-1} \mathbf{X}^T \mathbf{y}$$
(2)

## 2.2 Spatial Econometrics

Econometrics is a science that applies economics theory, economics mathematics, and economics statistics to provide empirical support from the models built by economics theory and to produce the results in figures. Economics theory plays a role to make hypotheses which are generally qualitative on the relationship between economics variables and modeled economics phenomena. Economics mathematics plays the role to formulate economics theory in the form of mathematical equations with no regard to theory empirically, while statistics plays the roles in collecting empirical data, changing a model into Econometrics model, data processing, selecting and applying precise estimates to regression parameters and testing hypotheses and model accuracy. An econometric model is considered

good when it is able to analyze structurally and to predict and evaluate policy (Setiawan and Kusrini, 2010).

The development of econometrics is econometrics spatial. From the development, the economic phenomenon occurred is presumed to have a relationship when viewed from the aspect of location and others. Beginning with Tobler's first law (Ho & Hensher, 2016), the conditions at one point or area correspond to the conditions at one point or the adjacent area.

Anselin described two spatial effects in econometrics including the effects of spatial dependence and spatial heterogeneity (Ciccarelli & Fachin, 2016). Spatial dependence indicates an autocorrelation among research object locations (cross sectional data set). Spatial heterogeneity refers to the diversity of functional forms and parameters at each location, the study sites did not indicate homogeneity in the data. In general, spatial model can be stated in the equation (3).

$$\begin{aligned} \mathbf{y} &= \rho \mathbf{W} \mathbf{y} + \mathbf{X} \boldsymbol{\beta} + \mathbf{u} \\ \mathbf{u} &= \lambda \mathbf{W} \mathbf{u} + \varepsilon \\ \varepsilon &\sim N \big( 0, \sigma^2 \mathbf{I} \big) \end{aligned} \tag{3}$$

where **y** is  $(N \ge 1)$  vector of the dependent variable, **X** is  $(N \ge p)$  matrix of the independent variable, **B** is  $(p \ge 1)$  vector in the regression coefficient, **W** is the  $(N \ge N)$  weight matrix which states the relationship between N,  $\rho$  and  $\lambda$ , which is the dependency spatial parameter, and  $\varepsilon$  is an independent and identical distributed vector (i.i.d).

#### 2.3 Weight Matrix

Spatial Weight Matrix (W) can be obtained from inter-regional contiguity and neighborhood distance or the distance between all regions. Spatial weights specifications vary according to conceptualizations; popular ones include: contiguity, continuous distance, nearest neighbor distances, fixed distance band, and spatial interaction. (Griffith, Griffith, & A., 2017)

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## 2.4 Spatial Durbin Error Model

(spatial error model) with the equation:

$$\mathbf{y} = \mathbf{X}\,\boldsymbol{\beta} + \mathbf{u} \tag{4}$$

 $u = \lambda W u + \varepsilon$ (5)

The equations of (4) and (5) can be completed and it can be obtained the value of u:

$$u = \lambda W u + \varepsilon$$
  

$$u - \lambda W u = \varepsilon$$
  

$$(I - \lambda W) u = \varepsilon$$
  

$$u = (I - \lambda W)^{-1} \varepsilon$$
(6)

From the equations of (4) and (6),

$$\mathbf{y} = \mathbf{X} \,\boldsymbol{\beta} + \left(\mathbf{I} - \boldsymbol{\lambda} \mathbf{W}\right)^{-1} \boldsymbol{\varepsilon} \tag{7}$$

Then, Lesage and Pace (2008) introduced spatial durbin error model (SDEM), with the addition of spatial lag in the independent variable

$$y_{i} = \beta_{0} + \sum_{k=1}^{l} \beta_{1k} x_{ki} + \sum_{k=1}^{l} \beta_{2k} \sum_{j=1}^{n} w_{ij} x_{kj} + (I - \lambda \mathbf{W})^{-1} \varepsilon$$
(8)

It can be rewritten in matrix notation as follows:

 $y = Z \beta + (I - \lambda W)^{-1} \varepsilon$ where  $\mathbf{Z} = \begin{bmatrix} \mathbf{I} \ \mathbf{X} \ \mathbf{W} \mathbf{X} \end{bmatrix}$  and  $\mathbf{\hat{a}} = \begin{bmatrix} \beta_0 \ \beta_1 \ \beta_2 \ \beta_3 \ \beta_4 \end{bmatrix}^T$ , WX are the spatial *lag* in X, and *I* is the identity matrix.

## 2.4 Estimation of the Parameter of Spatial Durbin Error Model (SDEM)

The method of Maximum Likelihood Estimation (MLE) was used to estimate the parameter of SDEM. From the equation (8), the likelihood function was formed, and the formation

of the likelihood function is made through error Lesage and Pace (2008) introduced the SEM  $\varepsilon$ . The result of the function formation is in the equation (9)

$$y = Z \beta + (I - \lambda W)^{-1} \varepsilon$$
  

$$\varepsilon = y (I - \lambda W) - (I - \lambda W) Z \beta$$
  

$$\varepsilon = (I - \lambda W) (y - Z \beta)$$
(9)

where,  $\mathbf{Z} = \begin{bmatrix} I X_1 X_2 & WX_1 & WX_2 \end{bmatrix}$ and  $\boldsymbol{\beta} = \left[ \mathsf{I} \, \beta_0 \beta_1 \beta_2 \beta_3 \beta_4 \beta_5 \right]$  $J = \left| \frac{\partial u}{\partial y} \right| = \left| I - \lambda \mathbf{W} \right|$ 

Then the result is;

$$L(\lambda,\beta,\sigma^{2};y) = (2\pi)^{-\frac{n}{2}} (\sigma^{2})^{-\frac{n}{2}} \left| J \right| e^{\left\{ -\frac{1}{2\sigma^{2}} \varepsilon^{T} \varepsilon \right\}}$$

$$L\left(\lambda,\beta,\sigma^{2};y\right) = \left(2\pi\right)^{-\frac{n}{2}} \left(\sigma^{2}\right)^{-\frac{n}{2}} \left|I - \lambda W\right| e^{\left\{-\frac{1}{2\sigma^{2}}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]^{T}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]\right\}}$$

Natural Logarithm Operation (In likelihood) in the equation of (16)

$$\ln L(\lambda,\beta,\sigma^{2};\mathbf{y}) = \mathbf{c} - \frac{\mathbf{n}}{2}(\sigma^{2}) + \ln \left| \mathbf{I} - \lambda \mathbf{W} \right| - \frac{1}{2\sigma^{2}} \left[ (\mathbf{I} - \lambda \mathbf{W})(\mathbf{y} - \mathbf{Z}\beta) \right]^{\mathrm{T}} \left[ (\mathbf{I} - \lambda \mathbf{W})(\mathbf{y} - \mathbf{Z}\beta) \right]^{\mathrm{T}}$$
(10)

From the equation of (10), the estimators of parameters  $\hat{\beta}$ ,  $\hat{\lambda}$  and  $\hat{\sigma}^2$  are obtained.

## **2.5** Estimation of Parameter $\hat{\beta}$

The estimator parameter  $\hat{\beta}$  was obtained by maximizing the function of ln likelihood of the equation (10), which is the first derivative of the equation to and makes it equal to zero, as follows:

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$$\frac{\partial L\pi \left(\lambda,\beta,\sigma^{2};y\right)}{\partial\beta} = 0$$

$$\frac{\partial \left\{c - \frac{n}{2}\ln(\sigma^{2}) + \ln\left|I - \lambda W\right| - \frac{1}{2\sigma^{2}}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]^{T}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]\right\}}{\partial\beta} = 0$$

$$\frac{1}{\sigma^{2}}\left\{\left[Z^{T}\left(I - \lambda W\right)^{T}\left(I - \lambda W\right)y\right] - \left[Z^{T}\left(I - \lambda W\right)^{T}\left(I - \lambda W\right)Z\right]\beta\right\} = 0$$

$$\beta \stackrel{\text{ff}}{=}\left[Z^{T}\left(I - \lambda W\right)^{T}\left(I - \lambda W\right)Z\right]^{-1}\left[Z^{T}\left(I - \lambda W\right)^{T}\left(I - \lambda W\right)y\right]$$
(11)

**2.6 Estimation of Parameter**  $\hat{\sigma}^2$  (10) to  $\hat{\sigma}^2$  and making it equal to zero, as The estimation of parameter  $\hat{\sigma}^2$  was follows: obtained by the first derivative of the equation

$$\frac{\partial L\pi \left(\lambda,\beta,\sigma^{2};y\right)}{\partial \sigma^{2}} = 0$$

$$\frac{\partial \left\{c - \frac{n}{2}\ln(\sigma^{2}) + \ln\left|I - \lambda W\right| - \frac{1}{2\sigma^{2}}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]^{T}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]\right\}}{\partial \sigma^{2}} = 0$$

$$-\frac{n}{2\sigma^{2}} + \frac{1}{2\sigma^{2}}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]^{T}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right] = 0$$

$$\sigma \mathcal{E} = \frac{1}{n}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]^{T}\left[\left(I - \lambda W\right)\left(y - Z\beta\right)\right]$$

#### **2.6 Estimation of Parameter** $\hat{\lambda}$

The estimator  $\hat{\lambda}$  cannot be obtained from the residual of OLS, and the estimator  $\hat{\lambda}$  is obtained from the explicit form of the concentrated ln likelihood function (Anselin, 2001). Substituting the equations (11) and (12) into the equation (10) and ignoring the constants, then:

$$\ln L(\lambda) = -\frac{n}{2} \ln \left\{ \frac{1}{n} (y - Z\beta)^{T} (y - Z\beta) \right\} + \ln \left| I - \lambda W \right|$$

Because it is not closed form, the completion to find the parameter estimation was conducted using iterative method.

## 2.7 Transport Infrastructure and Regional Economic Growth

(12)

In general, when viewed from the quality of infrastructure in Indonesia in the Report of the World Economic Forum 2017, Indonesia was in 41<sup>th</sup>rank out of 138 countries. Similarly, in 2016, although Indonesia's infrastructure development showed its progress in the 41<sup>st</sup> rank, it is still far left behind Singapore which ranked in the 2<sup>nd</sup>, Malaysian in 25<sup>th</sup> rank, and Thailand in 34<sup>th</sup> rank as outlined in Figure 1.

The Government through the Presidential Regulation No. 75 of 2014 on the Committee for the Acceleration of Infrastructure Provision describes several types of infrastructure provisions

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regulated by the government, i.e.: transport infrastructure, road infrastructure, irrigation infrastructure, drinking water and sanitation infrastructures, telematics infrastructure, electricity infrastructure, and oil transport and natural gas infrastructures. The classification of the above infrastructures can be categorized as basic infrastructures because they are needed by broader community and it has to be regulated by the government.



Figure 2 Achievement of Global Competitiveness Pillar in 2016 Source: Competitiveness Global Report 2016-2017

Transport infrastructure plays a role in creating the value of a good. Based on neoclassical theory, an item has the value according to the production cost or specifically by labor sacrifice cost issued thereon. Transportation is a tool that can create higher value in goods so that the goods can meet consumer needs (Li et al., 2017). In this case, transportation provides the value for goods through the process of moving the goods from the center of production to the center of consumption. The creation of the value for goods by transportation makes transportation a valuable tool economically.

Thus, the development of transportation infrastructure is one important aspect in accelerating the process of national development as well as driving sectoral economic growth.

## 3. METHOD

## 3.1 Data

The data used in this research were the data obtained from the Statistic Indonesia (BPS) of the Districts and Cities in Central Java Province in 2015.



Avalaible online at http://journals.ums.ac.id, Permalink/DOI: 10.23917/jep.v18i2.4660 Jurnal Ekonomi Pembangunan, 18 (2), 2017, 213-224 **Table 1 Definition of Operational Variables** Analysis Data No Variable Indicator Unit Sources Endogenous (Y) BPS Million GRDP (Y) 1 Gross Regional Domestic Product (GRDP) District/ Rupiah City Exogenous (X) Labor value for each district and city in Central BPS Number of Labors 2 Java person District/ (X2) City Human capital approached with the number of BPS 3 Human capital (X3) residents with the lowest education of Junior High District/ person School for each district and city in Central Java City BPS Length of roads with the conditions of good and Transportation 4 medium (km) for each district and city in Central Km District/ Infrastructure (X4) Java City

## 3.2 Research Variable and Spesifications Models

This study used the variables used by (Arbués et al., 2015) who used SDM with the Cobb-Douglas methodology. The difference was in the development of the spatial durbin error model. Table 1 shows the definitions of the variables used.

The designs of spatial econometrics model with two proposed spatial processes are:

a. SEM Model

$$y = \beta_0 + X_1\beta_1 + X_2\beta_2 + X_3\beta_3 + X_4\beta_4 + (I - \lambda W)^{-1}u$$

b. SDEM model:

 $y = \beta_0 + X_1\beta_1 + W_1X_1\beta_1 + X_2\beta_2 + W_2X_2\beta_2 + X_3\beta_3 + W_3X_3\beta_3 + X_4\beta_4 + W_4X_4\beta_4 + (I - \lambda W)^{-1}u$ 

## 3.3 Analysis Method

To model the response variables, the data structure processed can be seen in Table 2.

Table 2.	Data	Structure	of Rice	Production
Modelling				

District/City	Y	X <sub>1</sub>	$\mathbf{X}_2$	$\mathbf{X}_{3}$
1	$\mathbf{y}_1$	<b>x</b> <sub>11</sub>	$\mathbf{x}_{21}$	$\mathbf{x}_{31}$
2	$\mathbf{y}_2$	$\mathbf{x}_{12}$	$\mathbf{x}_{22}$	$\mathbf{x}_{32}$
÷	÷	÷	÷	:
35	$y_{35}^{}$	<b>X</b> <sub>135</sub>	$X_{235}$	X <sub>335</sub>

The followings are the stages of the analysis for each method in accordance with the formulation of research problems,

- a. Mapping the GRDP value of each District and City in Central Java as the initial study to determine the spatial weight matrix.
- b. Modelling the spatial error model (SEM), and spatial durbin error model (SDEM) using the method of maximum likelihood
- c. Evaluating the spatial econometrics model formed by testing the error assumption so that the system of assumption regression

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Parameter	OLS Coefficient (P-value)	SEM Coefficient (P-value)	SDEM Coefficient (P-value)
Intercept	8.005e-05	-0.011	0.44
	(0.999)	(0.916)	(0.025*)
Labor	1.814e-01	0.227	-0.007
	(0.446)	(0.294)	(2,2e-16*)
Human Capital	6.762 e- 01	0.646	-0.14
	(0.003*)	(0.001*)	(2,2e-16*)
Road Infrastructure	2.611e-01	-0.283	-0.040
	(0.137*)	(0.077**)	(0.068**)
W*Labor			0.244
			(0.014*)
W*Human Capital			-0.218
			(0.015*)
W*Road Infrastructure			-0.116
			(0.227)
Lambda	-	-0.019	-0.125
	-	(0.81)	(0.09**)
AIC	84.326	86.268	79.669

#### Table 3 Results of Parameter Estimation Test of OLS, SEM and SDEM Models

Note: \*Significant at  $\alpha = 5\%$ , \*\* Significant at  $\alpha = 10\%$ 

equation was obtained. The error assumptions in the spatial model included normal distributed errors, the error with a constant variance and free error.

#### 4. Result and Discussion

4.1 Modelling of Rice Production in the Districts/ Cities in Central Java using Spatial Econometrics Approach

The estimations of OLS, SEM and SDEM models resulted in the parameters influencing economic growth in Central Java province with a significance level of 5%. The result of the parameter estimation is presented in Table 3 as follows:

Based on the analysis on OLS, SEM and SDEM models, it can be concluded that the SDEM model is better to use because it has the smallest AIC value. SDEM is different from OLS and SEM models, because in SDEM there is a direct effect (global multiplier) and an indirect effect (multiplier). The indirect effects of SDEM are apparent from the parameters weighted by matrix W. All variables which are not weighted by matrix W (global multiplier) in SDEM have significant effect on GRDP. In other hand, the variables weighted by matrix W (local multiplier), the variables of labor and Human Capital have the significant effects on GRDP.

The lambda coefficient in SDEM is negative and significant at alpha 10 percent. It indicates that the GRDP in the adjacent districts and cities in Central Java are linked to each other, but in contrary. It means; if an area has a high economic growth, the areas nearby do not feel the significant impact of the economic growth. These coefficients are similar to those obtained in some of the most recent applied studies in Spain (Marquez, Ramajo, & Hewings, 2010). While the results of direct effect from Transportation Infrastructure and Labor does not match the model specification of negative marks similar to estimates by Cantos, Gumbau, Albert, & Maudos (2005). However,

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these two variables have no significant effect on GRDP Central Java province. This result can be interpreted that GRDP Central Java province is influenced by Human Capital globally.

Figure 4 shows the GRDP distribution of the districts and cities in Central Java in 2015. From the figure, it shows that there are regional groupings with high, medium and low GRDP values. The cities of Semarang, Cilacap and Kudus are included as the regions with high GRDP value, while the cities of Tegal, Magelang and Pekalongan are classified as the regions with low GRDP value. Figure 4 can show that the lambda value of GRDP is negative, which shows some regions with high GRDP value but the surrounding regions have low GRDP, such as Semarang and Demak, Demak and Kudus, Brebes and Tegal, Banyumas and Purbalingga, Semarang and Grobogan, and other areas showing a negative relationship.



Source: Jawa Tengah dalam angka 2016 (Central Java in Figures (processed) Figure4. The Distribution of GRDP of the Districts and Cities in Central Java `

## 4.2 Diagnostic of Models Assumption

Table 4 Evaluation of Model Assumption				
Classic Assumption Test	Criteria	SDEM		
Normality	p-value > $\alpha$	М		
Autocorrelation	p-value $> \alpha$	M*		
Heteroscesdacity	p-value > $\alpha$	$\mathbf{M}$		
Multicollinearity	VIF < 10	М		

Note: M = Meeting the econometrics criteria,

TM = Not meeting the econometrics criteria,

\*the autocorrelation test used global moran's test

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Table 4 is an evaluation of classic assumption results. The Kolmogorov-Smirnov test of SDEM result p-value of 0.1475 that indicates the data is normally distributed. Moran's results show p-value of 0.5059 which means that no autocorrelation occurs. Then, the Breusch-Pagan test obtained p-value of 0.066 and the VIF value for each variable has the value lower than 10. It means that SDEM has met the criteria of regression classic assumption.

## 5. Conclusion

The result of study have proposed an alternative modeling for economic growth in Central Java using spatial econometrics approach, we have performed several estimation procedures spatial modeling. First, we have estimated the model using ordinary least square (OLS) without spatial effect. Secondly, we have estimated two models spatial econometrics using spatial error model (SEM) and spatial durbin error model (SDEM). The results of our estimation show SDEM is the best model to model the economic growth of Central Java. SDEM shows that all variables that are not weighted by W have significant direct effect on the economic growth of Central Java, while the variables weighted by matrix W, Labor and Human Capital, and have the significant indirect effect on the economic growth of Central Java.

The empirical results have shown the presence of spatial effect in particular appears to support that distribution of economic growth in Central Java Province has a grouping pattern between the adjacent regions to each other. The role played by spatial effects has constituted the main focus of this analysis, the spatial effects are captured by indirect effects (spillovers).

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